

Gendered trends in income poverty in post-apartheid South Africa, 1997 – 2006

Dorrit Posel and Michael Rogan

posel@ukzn.ac.za and rogan@ukzn.ac.za

University of KwaZulu-Natal

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Abstract

This study investigates whether trends in the extent and depth of poverty in South Africa over the past decade have been gendered. We examine first, whether females are more likely to live in poor households than males, and whether this has changed over time; and second, how poverty has changed among female-headed and male-headed households. We use data provided by the 1997 and 1999 rounds of the October Household Survey and the 2004 and 2006 rounds of the General Household Survey. These surveys have the advantage of collecting information on the individual receipt of social grant income. We test whether our findings on gendered trends in poverty are robust to the possible underestimation of household income and to adjustments for household composition.

Keywords: gender, poverty and female-headed households

1. Introduction

Poverty studies from a number of countries and regions, spurred on by the United Nations Development Programme's (1995) claim that seventy per cent of the world's poor are women, have investigated gendered trends in poverty over the past decade. In South Africa, amidst an ongoing debate over the nature of trends in overall poverty since the end of apartheid, several studies have either hypothesised (cf. Phalane, 2002; Bentley, 2004; Borat et al., 2006; Thurlow, 2006) or offered preliminary evidence of (Bhorat and van der Westhuizen, 2008) a 'feminisation' of income poverty. Research has also documented a rise in female headship over the period (Bhorat et al., 2006), with female-headed households being over-represented among the poor (Posel, 1997; Leibbrandt and Woolard, 2001; Borat and van der Westhuizen, 2008). Against this backdrop, several authors have suggested that post-apartheid economic policy may not have been successful in redressing gender differences in the burden of poverty (Taylor, 1997; Phalane, 2002; Thurlow, 2006).

To date, however, there has been no comprehensive study of gendered poverty trends in the post-apartheid period. This may be explained partly by the limited availability of comparable surveys, conducted at regular intervals, which collect comprehensive information on household income or expenditure. In this paper, we analyse the income and expenditure data provided in two sets of nationally representative household surveys, which have been relatively under-utilised in poverty studies thus far: the 1997 and 1999 October Household Surveys and the 2004 and 2006 General Household Surveys. Although these surveys do not capture all sources of income received in households, they do consistently collect information on the individual receipt of both earned income and social grant income. We use these data to investigate whether trends in the extent and depth of poverty in South Africa over the past decade have been gendered. We also test the robustness of our findings to the possible underestimation of household income and to adjustments for household composition.

The paper is structured as follows. The next section summarises the existing body of literature on gender and poverty in South Africa. Section 3 discusses the data and methods that we use to investigate gendered trends in poverty. In Section 4 we present estimates of poverty trends by gender over the post-apartheid period and in Section 5 we evaluate the robustness of these trends. Section 6 concludes with a summary of the estimated changes in gendered poverty rates during the post-apartheid period.

2. Review

The post-apartheid period has brought a number of developments which are likely to have increased women's access to resources. Since 1995, for example: women's share of employment has grown (Casale and Posel, 2002; Casale and Posel, 2005); the social security system has been expanded to include support for pensioners and the caregivers

of children across all race groups; and progressive labour legislation, which includes the extension of minimum wages for domestic workers, has been introduced.¹ Despite these positive developments, there are other changes which suggest that women's relative access to resources may have worsened over the period. Gender analysts point to the disproportionate effects of the HIV/AIDS epidemic on women (Bentley, 2004; Schatz and Ogunmefun, 2007), rising rates of unemployment among women (Casale and Posel, 2002; Casale, 2004), and a widening gender gap in real earned income (Leibbrandt et al., 2005) together with the increasing proportion of informal and low-income workers who are women (Casale, 2004).

The literature on poverty in post-apartheid South Africa has paid little attention to the gendered nature of poverty. Rather, studies have focused on measuring aggregate trends in poverty and, given the legacy of apartheid, on changes disaggregated by race. Research suggests that until about 2000, overall income poverty rates remained relatively unchanged or even increased (Leibbrandt and Woolard, 2001; Borat and Kanbur, 2005; Hoogeveen and Özler, 2005; Leibbrandt et al., 2006; Seekings, 2007; Borat and van der Westhuizen, 2008; van der Berg et al., 2008). After 2000, and particularly in the period (post 2001) when government spending² on social grants increased and the take-up of the child support grant expanded, studies identify that poverty rates began to fall (Meth, 2006; Borat and van der Westhuizen, 2008; van der Berg et al., 2008). There remains some disagreement however over the extent of this decline in poverty (Meth, 2006; Seekings, 2007; van der Berg et al., 2008).

One of the key problems with measuring changes in poverty in the post-apartheid period is that comparable and comprehensive income or expenditure data have not been collected at regular intervals. The Population Censuses (1996 and 2001), conducted at long intervals, capture income in (inconsistent) bands and have been plagued by large numbers of zero-income households (Leibbrandt and Woolard, 2001). The Income and Expenditure Surveys (1995, 2000 and 2005) are also conducted at long intervals and changes in the survey methodologies³ between survey years compromise comparisons of the data (Statistics South Africa, 2008; van der Berg et al., 2008). The Labour Force Surveys (undertaken bi-annually from 2000 to 2007) provide a more regular source of data, but these surveys collect information on the individual receipt only of *earned* income.

¹ The 1997 Basic Conditions of Employment Act provided protection against unfair dismissal and enforced a minimum wage. The 1998 Employment Equity Act promotes the employment of both women and non-Whites across all sectors.

² Government expenditure on social grants increased in the post-2000 period (starting in 2001-2) with the proportion of GDP spent on social grants rising from about two per cent in 1994 to 3.5 per cent in 2005 (Seekings 2007).

³ The 2005 Income and Expenditure Survey uses a diary method to capture information on household expenditures over the past four weeks. In order to employ this new methodology, field workers visit each household five times over a four week period and leave an expenditure (or acquisition) diary with respondents over this period. This approach differs significantly from past surveys (1995 and 2000) where respondents were asked to recall their expenditures over the most recent four week period.

Measuring gendered changes in poverty brings the added difficulty of how to assign income across individuals in a household. In generating conventional estimates of income poverty, it is assumed that all household resources are equally shared among household members (or among adult equivalents). An individual is identified as being poor if he or she lives in a household in which average per capita (or per adult equivalent) income is below a poverty threshold. However, where resources are not equally shared among household members, then these poverty measures may conceal a gendered distribution of poverty within households. To explore the gendered nature of poverty further, studies from both developed and developing countries compare the economic well-being of female and male-headed households. If household heads are the primary income earners in households, then an analysis of poverty based on the gender of the household head highlights the nature and implications of gender differences in access to resources.

The poverty literature, particularly in developing countries, suggests that female-headed households are often more vulnerable to poverty. Out of 61 studies investigating the association between poverty and female-headed households in developing countries, summarised in Buvinic and Gupta (1997), 38 studies found that female-headed households were over-represented among poor households; 15 identified that poverty was associated with certain types of female-headed households or that, based on some poverty measures, a statistically significant relationship was found; and only eight studies reported no association between female headship and poverty. Similarly, in the World Bank's poverty assessments, poverty was higher for female-headed households than for male-headed households in 25 out of 58 countries. In a further ten countries, some types of female-headed households were poorer than male-headed households (Lampietti and Stalker, 2000).

Several South African studies have demonstrated that female-headed households are relatively disadvantaged in terms of income (Budlender, 1997; Posel, 2001) or are over-represented among the poor (Ray, 2000; Leibbrandt and Woolard, 2001; Bhorat and van der Westhuizen, 2008). However, there has been little research that has investigated how the relative well-being of female-headed households has changed over time. The one exception is a study by Bhorat and van der Westhuizen (2008), which uses expenditure data from the 1995 and 2005 Income and Expenditure Surveys. The study finds that although poverty incidence fell among both female and male-headed households, female-headed households became relatively more vulnerable to poverty over the period.⁴

⁴ The incidence of poverty among female-headed households decreased significantly from about 65.6 per cent to 60.6 per cent over the ten-year period. However, poverty rates among male-headed households declined from 45.8 per cent to 38.3 per cent, representing a greater absolute and relative fall. At the individual level, Bhorat and van der Westhuizen (2008) find that the proportion of poor individuals living in female-headed households increased from 42.3 per cent in 1995 to 54.8 in 2005.

3. Data and methods

In this paper we investigate gendered poverty trends in post-apartheid South Africa in two ways. First, we examine how poverty rates among females and males have changed over the period. Second, we estimate whether the economic well-being of female-headed households, relative to male-headed households, has changed.

We make use of data sources that have been under-utilised in analyses of poverty in the post-apartheid period. These include the 1997 and 1999 rounds of the October Household Survey (OHS) and the 2004 and 2006 rounds of the General Household Survey (GHS). Each of these surveys collects detailed information on earned income and importantly, also on the individual receipt of social grants. With the exception of the 1997 OHS which asks respondents for the annual value of social grant income, information on social grant receipt is collected only as a binary variable. For consistency across all the surveys, we derive a measure for social grant income in each survey by assigning the maximum value of the particular social grant to the grant recipient.

In using this approach to measure social grant income, our estimates of the extent of social grant spending are largely comparable to those documented in official administrative records.⁵ Estimates from the 2004 GHS, for example, place the total annual value of social grants at approximately R43 billion. National Treasury (2007) records show that the Department of Social Development spent a total of R44.8 billion on all social grants for the 2004/5 fiscal year. According to the 2006 GHS, total annual social grant spending is estimated at about R51.5 billion. In comparison, official records document spending at approximately R51.9 billion for the 2005/6 fiscal year and R56.9 billion in 2006/7 fiscal year (National Treasury, 2007).

One of the advantages of the household survey data used in this study is that we can isolate the relative contributions of social grant income to household resources, *ceteris paribus*, and evaluate the poverty-reducing effects of these transfers. Nonetheless, several problems remain. First, none of the surveys collects comprehensive information on private pensions. As a result, many of the elderly, living in households where neither social grants nor earned income is reported, would be incorrectly identified as being in poverty. Second, only the 1997 OHS collects information on the value of remittances and private maintenance, income sources that are more likely to be important for low-income households and for female-headed households (Leibbrandt and Woolard, 1999).

In Table 1, the first two data rows describe the number and proportion of households for which no earnings or social grants are reported in the datasets which we analyse. Although there is a considerable decline in these “zero-income” households across the years, almost 15 percent of all households in the GHS 2006 still report receiving neither earnings from employment nor grant income. To address this problem, we use

⁵ While overall spending on social grants captured through household surveys is in line with administrative records, there is a weaker correspondence in the composition of grants. For example, household survey data over-estimate the receipt of the social pension and considerably underestimate the receipt of the disability grant.

information on household expenditure to augment household income in “zero-income” households.⁶

Although household expenditure is captured only through a single question, and is therefore a relatively crude measure of economic status, it offers the means to approximate income, and particularly poverty status, in households which do not report earnings or grant income. Table 1 shows that using expenditure as a proxy for income in households eliminates zero-income households from the 1997 OHS, and in the following years, the proportion of zero-income households is reduced to 3.95 per cent, 1.34 per cent and .60 per cent of all households, respectively.⁷

Table 1: Zero income households, 1997 – 2006

	OHS 1997	OHS 1999	GHS 2004	GHS 2006
Zero earnings + zero social grants				
Number	6 972	6 273	4 449	4 031
Proportion	23.39	24.00	16.97	14.40
Including household expenditure				
Number	0	1 033	350	167
Proportion	0.00	3.95	1.34	0.60
Real per capita monthly household income (2000 prices)	279.12 (6.85)	210.70 (4.57)	443.66 (10.63)	434.13 (12.18)

Source: Own calculations from the 1997 and 1999 OHSs and the 2004 and 2006 GHSs

Notes: Zero income households are households that receive neither earnings nor social grants.

Real per capita income is only for zero-income households.

Data are weighted in the last row of the table

Standard errors are in brackets

In the poverty analysis, we estimate the overall incidence and depth of poverty and look specifically at how these have changed by gender and across female- and male-headed households over the specified period. The headcount and poverty gap ratios are measured using the normalised Foster-Greer-Thorbecke (FGT) index:

$$P_{\alpha} = 1/n \sum_{i=1}^m [(z - y_i) / z]^{\alpha}$$

where n represents the population size, m is the number of people falling below the poverty line z, and y_i represents the average per capita monthly household income of

⁶ The 1997 OHS captures household expenditure as a point value while the other surveys collect expenditure information in bands. Where expenditure is captured in bands, we assign the midpoint of the expenditure bracket to households that report zero earnings and zero social grant income. Most households with zero income did not report expenditure in the lowest band. In 1999, 2004 and 2006, between 34.5 per cent and 41.5 per cent reported that their household expenditure was in the lowest band (R0-R399 total household monthly income). Income bands are constant across all three surveys.

⁷ Across the four survey years, the real median difference between income and expenditure in households that report both income and expenditure information is positive and ranges from approximately R106 to R264.

individual i . The FGT index identifies the headcount or poverty rate when α takes the value of zero, and the depth of poverty when α equals one. We follow Hoogeveen and Özler (2005) in selecting R322 per capita monthly household income (in 2000 prices) as a plausible lower-bound poverty threshold for South African households. This also allows for comparability with a number of other poverty studies that have used this poverty line (or very similar poverty thresholds) to measure poverty in South Africa (cf. Hoogeveen and Özler, 2005; Ardington et al., 2006; Leibbrandt et al., 2006; Borat and van der Westhuizen, 2008).

In order to estimate the extent and depth of poverty, we consider three different measures of per capita monthly household income: earned income only (measure I); earned income and social grant income (measure II); and earned income and social grant income with household expenditure used as a proxy for income in zero-income households (measure III). The first measure of income identifies how poor individuals would have been had they relied only on the earnings of resident household members. The second measure helps highlight the poverty-reducing “effects” of social grant income. With the third measure of income, we offer some correction for the upwards bias in poverty estimates created by incomplete information on all income sources.

4. Poverty analysis

4.1 Individual poverty estimates

To compare poverty rates by gender, we first estimate poverty at the level of the individual. Individuals are poor if they live in households where average per capita monthly household income is below the poverty line of R322 (2000 prices). Because poorer individuals live in larger households in South Africa, poverty measured at the individual level is considerably higher than poverty measured at the household level (described in Section 4.2 below).

Table 2 reports headcount poverty rates ($\alpha = 0$) from 1997 to 2006, for the three measures of income. Within each year, poverty rates are lowered considerably as the measure of income becomes more comprehensive. In 1999, for example, 68.1 per cent of South Africans lived in earnings-poor households. With the receipt of social grant income, the headcount rate falls to 66.3 per cent; and when the income measure is augmented with data on household expenditure, the poverty rate declines further to 63.6 per cent.

Over the period 1997 to 2006, the poverty headcount decreased overall, but this fall masks an initial increase in poverty rates from 1997 to 1999. Poverty incidence then begins to decline between 1999 and 2004, with the decline most pronounced when both social grants (II) and household expenditure (III) are included as income sources (a significant decrease in the poverty headcount from 63.6 per cent to 61.6 per cent). From 2004 to 2006, there is a large and significant fall in poverty rates driven particularly by the poverty-reducing nature of social grant income. In 2006, approximately 64.3 per cent of all South Africans would have been poor had they relied only on income earned from

employment. With the inclusion of social transfers, however, the extent of poverty falls by five percentage points, to 59.1 per cent.

Table 2: Poverty estimates for South Africa, 1997 – 2006 (per capita)

	Headcount ($\alpha = 0$)			
	OHS 1997	OHS 1999	GHS 2004	GHS 2006
	Earnings only (I)			
All	65.33 (.006)	68.08* (.006)	68.87 (.007)	64.26* (.019)
Male	62.73 (.006)	65.61* (.006)	65.96 (.007)	60.77* (.019)
Female	67.74 (.006)	70.41* (.006)	71.69 (.007)	67.67* (.019)
	Earnings + social grants (II)			
All	62.46 (.006)	66.26* (.006)	65.25 (.007)	59.07* (.018)
Male	59.97 (.006)	63.94* (.006)	62.37* (.007)	55.70* (.018)
Female	64.78 (.006)	68.43* (.006)	68.05 (.007)	62.36* (.018)
	Including household expenditure (III)			
All	59.51 (.006)	63.62* (.006)	61.60* (.007)	55.96* (.020)
Male	57.11 (.007)	61.32* (.007)	58.40* (.007)	52.28* (.020)
Female	61.75 (.007)	65.78* (.007)	64.70 (.007)	59.55* (.020)

Source: Own calculations from the 1997 and 1999 OHSs and the 2004 and 2006 GHSs

* Denotes a significant change in the poverty estimate from the previous year at the 95% level of confidence. The data are weighted. Standard errors are in brackets. The poverty line is R322 per capita per month in 2000 prices. Household well-being is estimated as average per capita total household monthly income

The overall trend of an initial rise in the extent of poverty (1997-1999) followed by a fall in poverty rates (in 2004 and 2006) is consistent with trends identified in the recent poverty literature in South Africa (Bhorat and Kanbur, 2005; Leibbrandt et al., 2006; Meth, 2006; Seekings, 2007; Bhorat and van der Westhuizen, 2008; van der Berg et al., 2008). The actual poverty estimates that we present with the most comprehensive income measure (measure III in Table 2) are comparable to those identified by Ardington et al. (2006), who applied multiple imputation techniques to income data from the 1996 and 2001 Population Censuses to estimate income for households with missing and

implausible income values. Employing this technique, they find that the poverty headcount increased from 59.8 per cent to 65.1 per cent between 1996 and 2001.⁸

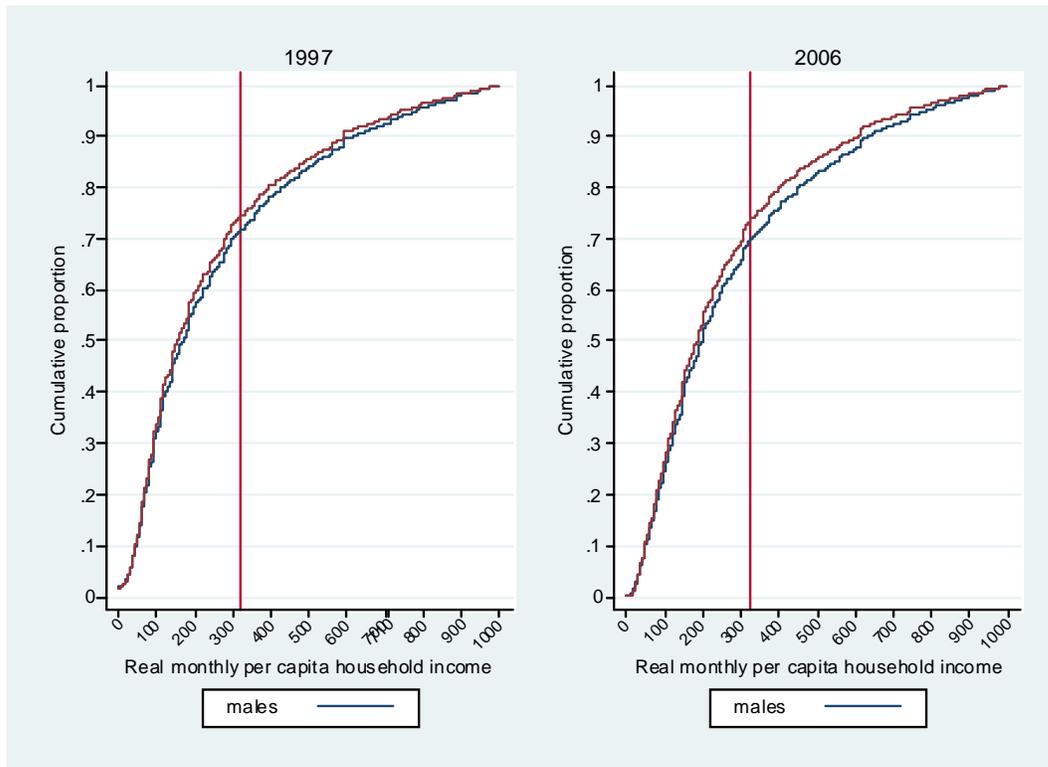
In Table 2, we further disaggregate poverty rates by gender. In each year, poverty estimates are significantly and consistently higher for females than for males across all three measures of income. In 1999, for example, estimating the poverty headcount using income measure III, 61.3 per cent of males lived in poor households compared to 65.8 per cent of females. Trends in poverty rates disaggregated by gender show a similar pattern to the overall trend, of headcounts rising for both males and females from 1997 to 1999 and then falling particularly from 2004 to 2006. However, the data also suggest that gender differences in poverty widened over the decade. In 1997, using the most comprehensive measure of income, approximately 57.1 per cent of males lived in poor households compared to 61.8 percent of females. By 2006, poverty rates had fallen to 52.3 per cent among males, but only to 59.6 per cent among females.

A clear visual representation of gender differences in poverty rates and of the widening gender gap between 1997 and 2006 is provided in Figure 1. The figure describes a cumulative distribution function (CDF) for males and for females living in households with average real per capita household monthly income of less than R1 000 per month (2000 prices) in 1997 and 2006, with the vertical line denoting the poverty line of R322.⁹ In both 1997 and 2006, the CDF for males falls below that for females at any point in the income distribution (indicating first order dominance). Furthermore, the distance between the CDFs for male and female income is noticeably wider in 2006 than in 1997. The figure therefore shows that at any plausible poverty line (both above and below R322) the gap in real monthly incomes between males and females has widened over the period.

⁸ Ardington et al. (2006) use a poverty line of R340 in 2001 prices (which translates into R322 in 2000 prices).

⁹ We have restricted the size of the population to aid the visual representation, but in so doing, the proportion of the population that is shown as being in poverty is higher than the national measures.

Figure 1 Cumulative distribution function of real monthly per capita household income by gender, 1997 and 2006



Source: Own calculations from the 1997 OHS and the 2006 GHS

The poverty gap ratios presented in Table 3 suggest the contribution of social grant income particularly to reducing the depth of poverty. In 1997, for example, the poverty gap ratio drops from .53 on the basis of earned income only, to .44 with the inclusion of social grant income; and in 2006 it drops also from .53 but to .36. Furthermore, the table shows that although poverty gap ratios are consistently higher for females than for males, the receipt of grant income reduces particularly the depth of female poverty. Across all four years, the greatest absolute change in the poverty gap after including social grant income occurs among females. In 2006, for example, the poverty gap declines from .50 to .35 for males, once grant income is included, but it drops from .56 to .38 for females.

Table 3: The poverty gap, 1997 – 2006

Poverty Gap ($\alpha = 1$)				
	OHS 1997	OHS 1999	GHS 2004	GHS 2006
Earnings only (I)				
All	.53 (.006)	.56* (.006)	.58* (.006)	.53* (.019)
Male	.50 (.006)	.54* (.006)	.55 (.007)	.50* (.019)
Female	.55 (.006)	.58* (.006)	.60* (.007)	.56* (.019)
Earnings + social grants (II)				
All	.44 (.005)	.49* (.005)	.42* (.005)	.36* (.011)
Male	.42 (.005)	.47* (.005)	.41* (.005)	.35* (.012)
Female	.46 (.005)	.50* (.006)	.43* (.005)	.38* (.011)
Including household expenditure (III)				
All	.36 (.005)	.41* (.005)	.36* (.005)	.31* (.013)
Male	.34 (.005)	.39* (.005)	.34* (.005)	.29* (.013)
Female	.37 (.005)	.42* (.005)	.37* (.005)	.33* (.013)

Source: Own calculations from the 1997 and 1999 OHSs and the 2004 and 2006 GHSs

* Denotes a significant change in the poverty estimate from the previous year at the 95% level of confidence. The data are weighted. Standard errors are in brackets.

To explore further the contribution of social grant income to poverty alleviation, we decompose the FGT measures of poverty using the Shapley-value (Shorrocks, 1982; Duclos and Araar, 2006). This method makes it possible to identify the average marginal effect of different income components to poverty reduction while recognising the different distributions of income components across individuals. Rather than simply calculating a share composition of each income source, the Shapley-value measures the average marginal effect of including an income source over all possible combinations of income sources for the population (Duclos and Araar, 2006). With the available data, we distinguish between three components of income: earned income; social grant income; and “other income” accruing to households which received neither earnings nor grant income.

Table 4 reports the average marginal contributions of each income source to reducing the incidence and the depth of poverty in 1997 and 2006. Income from employment has the largest marginal effect on both the rate and depth of poverty. From 1997 to 2006, the average relative contributions of earned income fell, however, while those from grant income increased. The contribution of grant income to poverty reduction derives particularly from its relative contribution to reducing the depth of poverty, and over the period this average marginal effect increased. In 1997 and 2006, grant income reduced

the poverty gap ratio by about 16 per cent and 26 per cent respectively. The table also highlights that grant income reduces the poverty rate, and particularly the poverty gap ratio, by more among females than among males.

Table 4. Relative contribution of income sources to reducing poverty

	All		Female		Male	
	1997	2006	1997	2006	1997	2006
Headcount ratio ($\alpha = 0$)						
Earned income	.88	.84	.86	.83	.88	.85
Grant income	.05	.09	.06	.10	.05	.08
Other income	.07	.07	.08	.07	.07	.07
Total	1.00	1.00	1.00	1.00	1.00	1.00
Poverty gap ratio ($\alpha = 1$)						
Earned income	.71	.66	.69	.64	.73	.68
Grant income	.16	.26	.17	.29	.15	.23
Other income	.13	.08	.14	.07	.12	.08
Total	1.00	1.00	1.00	1.00	1.00	1.00

Source: Own calculations from the 1997 OHS and the 2006 GHS using the DASP module developed by Araar and Duclos (2007).

Notes: The data are weighted

Overall then, the data presented in this section suggest that females are more likely than males to live in households where average per capita monthly household income lies below a poverty line of R322 (2000 prices). Consistent with other studies, we find that poverty rates have fallen, particularly over the 2000's. We find further that the incidence of poverty has declined for both males and females but that the fall has been absolutely and relatively smaller for females. Consequently, over the period 1997 to 2006, gender differences in the incidence of poverty have widened slightly in favour of males. With the receipt of social grant income in households, both the extent and depth of poverty are significantly lower than they would have been had households relied only on income earned through employment. These transfers are effective particularly in narrowing gender differentials in the depth of poverty.

4.2 Poverty estimates at the household level

To investigate gender trends in poverty further, we compare the incidence and depth of poverty in female and male-headed households. Although the value of self-reported headship as an analytical category is contested (cf. Budlender, 2005), evidence suggests that, in the majority of households in South Africa, household heads are also the main income providers. In the 1997 OHS, for example, approximately 83 per cent of all household heads earned or received the highest level of income in the household.¹⁰ Poverty estimates by the gender of the household head therefore offer an opportunity to explore the nature and the implications of gender differences in access to resources.

¹⁰ Across all four surveys, between 76 per cent and 83 per cent of household heads earned or received the highest level of income in the household.

Table 5 documents a steady increase from 1997 to 2006 in the number of households headed by women. This increase has been slightly larger than the rise in the number of households overall and consequently the proportion of all households which are female-headed grew modestly over the period, from 35.17 per cent to 37.48 per cent of all households.

Table 5: Male- and female-headed households, 1997 - 2006

	OHS 1997	OHS 1999	GHS 2004	GHS 2006
Male-headed	5 981 957 (35 144)	6 647 800 (43 464)	7 664 456 (65 833)	8 073 892 (82 953)
Female-headed	3 244 538 (25 996)	3 735 295 (35 973)	4 520 349 (49 458)	4 858 648 (57 856)
Proportion female-headed	35.17 (.003)	35.96 (.003)	37.07 (.004)	37.48 (.004)

Source: Own calculations from the 1997 and 1999 OHSs and the 2004 and 2006 GHSs
The data are weighted. Standard errors in brackets

In Table 6 we estimate headcount and poverty gap ratios using the most comprehensive measure of income (measure III), but now at the household level (households are poor if real average monthly per capita income in the household lies below the poverty line), and we distinguish among households headed by females and by males. Across all households, both the incidence and depth of poverty increased between 1997 and 1999 and then decreased in 2004 and particularly in 2006.

Table 6: Poverty estimates among households (measure III), by household type, 1997 – 2006

	OHS 1997	OHS 1999	GHS 2004	GHS 2006
Headcount ($\alpha = 0$)				
All households	48.63 (.007)	50.88* (.006)	48.23* (.007)	43.59* (.018)
Male-headed	38.84 (.007)	41.69* (.007)	37.24* (.007)	32.54* (.015)
Female-headed	66.68 (.008)	67.24* (.008)	66.88 (.008)	61.98* (.019)
Poverty gap ($\alpha = 1$)				
All households	.28 (.004)	.30* (.004)	.27* (.005)	.23* (.010)
Male-headed	.21 (.004)	.24* (.004)	.20* (.004)	.16* (.008)
Female-headed	.40 (.005)	.41* (.004)	.38* (.006)	.34* (.012)

Source: Own calculations from the 1997 and 1999 OHSs and the 2004 and 2006 GHSs
* Denotes a significant change in the poverty estimate from the previous year at the 95% level of confidence. The data are weighted. Standard errors are in brackets

In all years, female-headed households are far more likely to be poor and to lie further from the poverty line than male-headed households. Furthermore, although the headcount and poverty gap ratios have fallen among both male-headed and female-headed households, the relative decreases have been larger among male-headed households. For example, from 1997 to 2006, the headcount rate fell from 38.8 to 32.5 per cent (a fall of 16.2 per cent) among male-headed households but only from 66.7 to 62 per cent (a fall of seven per cent) among female-headed households.¹¹

To compare the relative contributions of the different income sources to poverty reduction in female- and male-headed households, we again make use of the Shapley-value decomposition. In Table 7 we compare the average marginal contributions of earned income, grant income and other income received in “zero-income” households to poverty reduction in 1997 and 2006. In both female- and male-headed households, income from employment had the largest contribution to reducing both the extent and depth of poverty. However, earned income is relatively more important in male-headed households.

Table 7. Relative contribution of income sources to reducing poverty

	All		Female-headed		Male-headed	
	1997	2006	1997	2006	1997	2006
Headcount ratio ($\alpha = 0$)						
Earned income	.83	.80	.71	.72	.86	.83
Grant income	.06	.10	.12	.16	.05	.08
Other income	.11	.10	.17	.12	.09	.09
Total	1.00	1.00	1.00	1.00	1.00	1.00
Poverty gap ratio ($\alpha = 1$)						
Earned income	.70	.67	.51	.52	.78	.74
Grant income	.14	.21	.24	.35	.11	.15
Other income	.16	.12	.25	.12	.11	.11
Total	1.00	1.00	1.00	1.00	1.00	1.00

Source: Own calculations from the 1997 OHS and the 2006 GHS using the DASP module developed by Araar and Duclos (2007).

Notes: The data are weighted

The marginal impact of grant income increased over the period: at the mean, the receipt of social grants reduced poverty rates in all households by six per cent in 1997 and 10 per cent in 2006. The table shows also that grant income is relatively more important in reducing the extent and depth of poverty in female-headed households. Although the expansion of the social grant system has not narrowed differences in poverty incidence between female- and male-headed households, it has lessened the depth of poverty by relatively more among female-headed households. In 2006, for example, social grant

¹¹ Male-headed households experienced a greater absolute and relative increase in their poverty headcount between 1997 and 1999, but this was more than offset, once poverty rates began to decrease (in 2004 and again in 2006), by a larger absolute and relative fall in poverty incidence.

income lowered the poverty gap ratio by 35 per cent in female-headed households compared to only 15 per cent in male-headed households. At the same time, however, the relative contribution of other income declined substantially in female-headed households that received neither income from employment nor from social grants. This may suggest that the receipt and value of remittances, particularly in households headed by women, has fallen over time.

The data presented in this section have demonstrated that poverty trends at the household level have also been gendered during the period under review. Both the poverty headcount and the depth of poverty have declined by relatively more among male-headed households compared with female-headed households. The decomposition analysis has suggested that, while social grant income was not enough to erase the poverty gap between male- and female-headed households, this income was relatively more important in reducing both the extent and depth of poverty among female-headed households. Differences in the incidence and severity of poverty between male- and female-headed households therefore would have been even greater without the receipt of social grant income.

5. Sensitivity tests

5.1 The under-estimation of income

One of the concerns with poverty studies that make use of household survey data is that income may be under-estimated, and poverty therefore over-estimated, because all sources of income earned or received in households may not have been measured. In the four surveys analysed in this study, information is consistently collected on the individual receipt of earnings and social grants. However, this is not the case for other sources of non-labour income and income transfers. As we illustrated earlier in Table 1, this resulted in a significant proportion of “zero-income” households, or households in which neither earned income nor social grant income was reported. In our poverty analysis, we addressed this problem by using data on reported household expenditure. In this section, we test whether this simple correction is an effective means of reducing the underestimation of household income.

Of the four datasets analysed in this study, only the 1997 OHS collects detailed information on the value of individual sources of income other than earnings and social grants. For the purposes of comparability across the years, we did not include this information in the poverty estimates for 1997 reported in the previous section. Here, we generate a new measure of household income (measure IV) for 1997, which includes the value of remittances received as well as of private maintenance, work pensions, unemployment benefits and gratuities. In Table 8 we present poverty statistics based on this new estimate of income and for ease of comparison, we repeat the statistics derived from the expenditure-augmented measure of household income (measure III) reported in earlier tables.

Table 8: Including other sources of income: Poverty estimates in 1997

	III	IV
Headcount (P₀)		
All individuals	59.51 (.006)	60.21 (.006)
Male	57.11 (.007)	57.90 (.006)
Female	61.75 (.006)	62.36 (.006)
All households	48.63 (.007)	50.67* (.006)
Male-headed	38.84 (.007)	41.34* (.006)
Female-headed	66.68 (.008)	67.89 (.007)
Poverty Gap (P₁)		
All individuals	.36 (.005)	.38* (.005)
Male	.34 (.005)	.37* (.005)
Female	.37 (.005)	.40* (.005)
All households	.28 (.004)	.32* (.004)
Male-headed	.21 (.004)	.26* (.005)
Female-headed	.40 (.005)	.44* (.006)

Source: Own calculations from the 1997 OHS

* Denotes a significant change in the poverty estimate from the preceding column. The data are weighted. Standard errors are in brackets. Measure III includes earned income and social grants, augmented with expenditure for zero-income households. Measure IV includes earned income and social grants plus remittance transfers, work pensions, private maintenance, unemployment benefits and gratuities.

The table shows that although differences in poverty estimates based on the two measures of income are not large, both the incidence and depth of poverty are *lower* using measure III than measure IV, and particularly among male-headed households.¹² One possible explanation for this is that even in the 1997 OHS, comprehensive information is not collected on all sources of income including private pensions and rental income, and measure IV therefore continues to underestimate income.¹³ These findings suggest that

¹² Among households which report information on both income and expenditure in 1997, the real median difference between per capita household monthly income and per capita monthly household expenditure is R138.23. The lower poverty estimates derived from measure III therefore are not likely to be caused by the overestimation of income from dissaving among households in 1997.

¹³ We also assessed whether our poverty estimates would have changed significantly if we had first included these other sources of income in measuring income in 1997, and then augmented income with data

using household expenditure to augment measures of reported income may be an effective means of redressing the underestimation of income in household surveys where information on all sources of income is not captured.

5.2 Adjusting for household composition

Following both international (cf. Ravallion, 1994) and South African (cf. Leibbrandt and Woolard, 1999) best practice, we test whether our gendered poverty estimates are robust to adjustments for household composition. In particular, we adjust both the poverty line and household income by the number of adult equivalents and by scale economies in the household (see United Nations Statistics Division, 2005). In choosing the most appropriate values for the adult equivalence and scale economies adjustments for South African households, we follow May et al. (1995) and Leibbrandt and Woolard (1999), assuming a scale parameter of .9 and that children consume half the resources of adults, and like Leibbrandt and Woolard (1999), we define children as those aged 10 and younger.¹⁴

Table 9 compares per capita estimates (measure III) of the extent and depth of poverty with estimates adjusted for household composition in 2006. The comparison suggests that estimating poverty rates using equivalence measures makes very little difference to overall poverty statistics, a finding which mirrors those reported in Leibbrandt and Woolard (1999) and Woolard and Leibbrandt (2001).

Table 9: Per capita and per adult equivalent poverty estimates (measure III), 2006

	Per capita (z=R322)		Per adult equivalent (z=R417) ¹⁵	
	P ₀	P ₁	P ₀	P ₁
All	55.96 (.020)	.31 (.013)	55.13 (.019)	.30 (.012)
Male	52.28 (.020)	.29 (.013)	51.83 (.019)	.28 (.012)
Female	59.55 (.020)	.33 (.013)	58.35 (.020)	.32 (.012)

Source: Own calculations from the 2006 GHS

The data are weighted. Standard errors are in brackets. No per adult equivalent poverty estimates are significantly different from the per capita estimates.

on household expenditure data. However we found very little difference with poverty estimates derived using measure III.

¹⁴. We therefore divide total household income not by the number of individuals but by $(A + .5K)^9$, where A is the number of adults and K is the number of children in the household.

¹⁵ We adjusted the R322 per capita monthly household income poverty line (2000 prices) for adult equivalence and economies of scale according to the method suggested by the United Nations Statistics Division (2005). We first identified a household of average size and composition in which per capita income equalled the unadjusted poverty line of R322. We then set the adjusted poverty line equal to household income adjusted for the number of adult equivalents and for the scale parameter.

In Table 10, we compare per capita and per adult equivalent trends in the extent of poverty using the most comprehensive measure of income (measure III). The table demonstrates that aggregate poverty trends and trends disaggregated by gender are not greatly affected by the use of an adult equivalence measure of income. Decreases in poverty rates among males and females are slightly greater when adjusting for household composition, but poverty rates among males continue to fall by absolutely and relatively more than poverty rates among females. Gendered trends in poverty are therefore robust to adult equivalent and economy of scale adjustments to household income.

Table 10: Per capita and per adult equivalent poverty trends

	Headcount (P_0)			
	OHS 1997	OHS 1999	GHS 2004	GHS 2006
	Measure III- per capita ($z= R322$)			
All	59.51 (.006)	63.62* (.006)	61.60* (.007)	55.96* (.020)
Male	57.11 (.007)	61.32* (.007)	58.40* (.007)	52.28* (.020)
Female	61.75 (.007)	65.78* (.007)	64.70 (.007)	59.55* (.020)
	Measure III- per adult equivalent ($z=R417$)			
All	58.60 (.006)	63.26* (.006)	60.86* (.007)	55.13* (.019)
Male	56.37 (.007)	61.04* (.007)	57.87* (.007)	51.83* (.019)
Female	60.68 (.006)	65.35* (.007)	63.74 (.007)	58.35* (.020)

Source: Own calculations from the 1997 and 1999 OHSs and the 2004 and 2006 GHSs

* Denotes a significant change in the poverty estimate from the previous year at the 95% level of confidence. The data are weighted. Standard errors are in brackets.

6. Conclusion

This paper has used income data from regularly collected household surveys to investigate gendered trends in poverty over the post-apartheid period. The advantage of using these datasets is that they include information on earned income and social grant income at the individual level. We compared poverty rates when individuals and households relied only on earned income with estimates that include social grants in the measure of income. In order to address the possible underestimation of reported income in households that received neither earnings nor social grants, we also used household expenditure as a proxy for income. To compare the relative contributions of the different income sources to poverty reduction, we used a Shapley-value decomposition.

The results presented suggest that, first and foremost, poverty remains a gendered phenomenon in post-apartheid South Africa. Across all four years and for each measure of income, the extent and depth of poverty is significantly higher for females and for

female-headed households. While the decline in overall poverty rates between 1997 and 2006 is encouraging, it masks differences in gendered poverty trends over the period. A greater decline in both the extent and the depth of poverty among males and male-headed households has resulted in a widening of the differences in poverty rates by gender. Particularly in the period during which declines in poverty have been greatest (2004-2006), the reduction in poverty has favoured males and male-headed households. From a gender perspective, however, the news is not all bad. Our findings suggest that the receipt of social grant income may have been relatively more effective in reducing particularly the depth of poverty among females and among female-headed households.

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